

Who Gets to Use Their College Skills? Skill-Job Alignment and Earnings Inequality Among College Graduates

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Abstract

College graduates who enter occupations closely linked to their majors – for example, engineering majors who become engineers – receive substantial earnings premiums. But many four-year degrees are not linked to a single profession, and labor markets increasingly reward transferable analytical, communicative, and quantitative skills shared across fields of study. I argue that skill transferability broadens the range of jobs in which graduates can realize the value of their human capital while opening a distinct margin of inequality in the college-to-work transition. Using linked course-catalog text, administrative transcripts, and employment records from a U.S. state, I develop a computational framework to measure transferable human capital and the alignment between graduates' coursework-derived skill profiles and occupational task demands. I find that this alignment varies substantially within majors and occupations. Graduates whose skills are more closely aligned with their work earn significantly more, even within the same occupation, but access to skill-aligned work is sharply stratified by race and class background: disparities in access account for a quarter of the Black-White earnings gap, net of observed human capital differences and sorting into higher- and lower-paying occupations. Much of higher education's economic

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value – and inequality in access to that value – depends on whether transferable skills are successfully matched to work across occupational boundaries.

1 Introduction

A college degree remains one of the most important routes to economic opportunity in the United States, but its rewards have become increasingly unequal. Even as the college wage premium stagnated in recent decades, the hourly wage gap between the highest- and lowest-paid college graduates grew by more than 50 percent (EPI, 2018). Against this backdrop, renewed debate over the value of a college degree has focused attention on whether and how college prepares students for the workforce. Public and scholarly debate has centered especially on the labor-market viability of particular fields of study, such as the humanities, and of lower-return institutions, including many in the for-profit sector (Bowen and Bok, 1998; Brand and Halaby, 2006; Witteveen and Attewell, 2017a; Gerber and Cheung, 2008; Altonji et al., 2014; Han et al., 2023; Kizilcec et al., 2023). While these literatures largely locate variation in economic returns to college in the educational investments themselves, a longstanding theoretical tradition recognizes that the payoff to educational investments depends on the labor-market settings in which acquired skills are deployed (Becker, 1962; Sattinger, 1993). To what extent, then, do graduates find jobs that use their college skills and with what consequences for inequality?

Despite the significance of this question, this dimension of the school-to-work transition remains poorly understood. Recent research instead focuses on a narrower issue: whether graduates enter occupations that are common for their field of study (Kalleberg, 2008; Witte and Kalleberg, 1995; de Werfhorst, 2002; Wolbers, 2003; Robst, 2007). This work has shown that workers often receive substantial earnings premiums when they follow established credential-to-occupation pathways, particularly in fields that channel students toward a small number of occupational destinations (DiPrete et al., 2017; Bol et al., 2019). These pathways provide an important mechanism for converting education into work, but they capture alignment most clearly when training is occupation-specific. Yet many four-year degrees in the United States do not impart occupation-specific skills, instead preparing graduates for a broad set of possible occupations. Contemporary employers likewise demand not only occupation-specific expertise but also transferable analytical, quantitative, communicative, and problem-solving capacities that can be deployed across a range of jobs (AACU, 2010; Autor et al., 2003; Moss-Pech, 2025). These features of higher education and work suggest that credential-to-occupation pathways capture only one way in which graduates may

put their education to use, necessitating a broader account of the relationship between what students learn in college and what they do at work.

This paper develops such an account. A major obstacle to studying this broader form of alignment is measurement: existing approaches use educational categories to proxy for what workers know. To overcome this obstacle, I develop a computational framework for measuring transferable human capital. Just as workers can transfer skills across occupations that rely on similar tasks (Gathmann and Schönberg, 2010; Yamaguchi, 2012), I argue that education can be conceptualized as developing task-relevant capacities that may transfer across jobs. I use the term *skill-task match* to describe the alignment between the skills graduates develop through coursework and the tasks they are required to perform at work.¹

I argue that transferable college-acquired skills can retain economic value across multiple labor-market domains, while also creating greater scope for inequality in whether graduates secure work that uses them. Unlike occupation-specific preparation, for which established credential-to-occupation pathways often identify plausible labor-market destinations, transferable skills may be valuable across a much wider range of jobs. A sociology graduate, for example, may acquire statistical, research, writing, and organizational-analysis skills that are useful in many different occupations. Realizing the value of those skills therefore depends on graduates' ability to identify and gain access to skill-aligned work. Because these capacities are unequally distributed, skill transferability may expand the range of valuable labor-market destinations while also generating inequality in who successfully reaches them.

To test this, I assemble a novel panel dataset that links administrative postsecondary transcripts to employment records for graduates in a U.S. state, and merge this with textual course-catalog data. The resulting dataset is unique in its ability to uncover the education-to-work transition at the level at which skill acquisition and deployment actually occurs. Rather than observing only declared majors, I use course-catalog text to infer the task-relevant content of each course a student takes in college, and measure skill-task match in terms of the similarity between a graduate's transcript-derived skill profile (aggregated across courses) and the task content of their post-college occupation.

I document three findings. First, there is substantial variation in whether graduates use their college skills in the U.S. labor market. Even among graduates from the same college and major and among graduates who enter the same occupation, some work in positions that are much more closely aligned with their college-acquired skills than others. Second, graduates who enter more skill-aligned work earn significantly

¹The framework is not specific to BA degrees. In principle, skill-task match could be measured for any postsecondary credential. In this paper, I focus empirically on BA graduates for reasons of measurement feasibility, as discussed in the Data section.

more than their peers, even within the same occupation. Third, access to skill-aligned work is highly stratified: Hispanic, lower-income, and Black graduates especially are significantly less likely to enter work aligned with their college skills than their more advantaged peers. These differences are economically consequential; disparities in skill-task alignment account for roughly one-quarter of the Black–White earnings gap, even net of differences in access to high-status occupations. Together, the findings identify skill-job alignment as an important and previously overlooked margin of earnings inequality among college graduates.

This paper shows that access to skill-aligned work beyond occupation-specific pathways is a motor of earnings inequality among college graduates. In so doing, it traces the roots of unequal returns to college past the well-documented educational inputs that shape earnings – such as college quality, field of study, and course-taking – and beyond sorting across the occupational hierarchy, to the relational position between graduates’ skills and their work. As technological transformations continue apace in altering the task structure of work, and as labor market demand for new combinations of analytic and technical capacities replaces old skill demands and coordination between education and work, the ability to transfer skills into new jobs and occupations may increasingly determine who benefits – and who loses out – from these transformations.

2 Education–Job Match and Credential Pathways

A hallmark feature of the contemporary knowledge economy is employers’ reliance on workers who possess specialized forms of knowledge, skill, and expertise. Some of these capacities can be developed through on-the-job training, and some are inherently specific to particular firms and organizational settings (Becker, 1962; Groysberg et al., 2008; Groysberg, 2010). But firms also rely on human capital developed outside the workplace. The educational system is an important source of this preparation, equipping workers with capacities they carry into the labor market and can deploy across jobs and employers. Such preparation becomes increasingly important as employers expect workers to enter the workplace already equipped with the requisite expertise (Cappelli, 1999, 2015), rather than selecting primarily on general abilities and developing role-specific skills through training and internal labor markets (Jacoby, 1983).

The pathway from schooling to work therefore poses an important matching problem: workers seek jobs well-suited to their college-acquired expertise, while employers seek workers who are sufficiently prepared to perform the tasks required for open positions. For the worker, the economic value of educational investments consequently depends not only on what they learn, but on whether that learning can be adequately deployed at work (Jovanovic, 1979; Sattinger, 1993). The same set of capacities can

be highly productive in one labor-market setting and comparatively underutilized in another.

Sociological research on education-job mismatch has developed this relational insight in two ways. Early work examined *vertical mismatch*: whether workers possess more or less schooling than their jobs require, documenting penalties associated with overeducation and underutilized schooling (Kalleberg, 2008; Leuven and Oosterbeek, 2011). But while vertical mismatch captures an important form of failed translation between schooling and work, workers with the same amount of education can possess very different capacities, and jobs requiring the same educational level can demand very different kinds of work. A second tradition has therefore examined *horizontal match*: whether the specific content of workers' education aligns with the work they enter (de Werfhorst, 2002; Wolbers, 2003; Robst, 2007).

Because the skill content of schooling is difficult to observe directly, research on horizontal match has generally operationalized alignment through credentials and occupational destinations. Early studies, using researcher-defined classifications of aligned occupations, found that graduates often earn more when they work in jobs connected to their educational field (de Werfhorst, 2002; Wolbers, 2003; Bol et al., 2019), although the size of the match premium has been disputed (Bédoué and Giret, 2011). More recent work has extended this approach by examining the payoff to following an occupational pathway that is common or expected given a worker's credential, and asks how the strength of school-to-work linkages moderates the payoff to doing so (Roksa and Levey, 2010; DiPrete et al., 2017; Forster and Bol, 2018; Bol et al., 2019). This literature shows that following a common credential-to-occupation pathway is particularly consequential in strongly linked fields, where programs channel graduates toward a relatively narrow set of occupational destinations.

These common pathways identify an important *coordinated* form of education-work alignment. In tightly linked fields, educational curricula, occupational categories, recruitment practices, professional institutions, and sometimes licensing rules jointly organize the movement from school into work. A nursing credential, for example, both indicates a relatively specific body of preparation and directs graduates toward a recognizable set of positions in which that preparation is expected to be useful. The credential makes graduates' capabilities legible to employers, while institutionalized pathways tell graduates where those capabilities belong. Under these conditions, observing that a graduate follows a common field-to-occupation pathway is a plausible proxy for the deployment of educationally acquired skills.

But many four-year degree programs organize the relationship between education and work differently. Fields such as the humanities, social sciences, and many natural sciences often prepare graduates for a range of possible destinations rather than a single expected occupation. More broadly, the United States has a comparatively weakly

coordinated system of skill formation: relative to countries with stronger vocational institutions, educational credentials are less tightly standardized around particular occupations and school-to-work linkages are weaker (Hall and Soskice, 2001; Estévez-Abe et al., 2001; Thelen, 2004; DiPrete et al., 2017). Liberal-arts curricula further allow students within the same major to complete different course sequences and acquire different combinations of skills (Han et al., 2023; Kizilcec et al., 2023). In these settings, field of study provides a weaker guide to both what graduates have learned and where those capacities can be productively deployed.

At the same time, many capacities demanded in the contemporary labor market are not the purview of a single major or professional program. Employers increasingly value nonroutine analytic skills such as quantitative, analytical, interactive, and communication-intensive capacities alongside more explicitly vocational preparation (Autor et al., 2003; Deming and Kahn, 2018; AACU, 2010). Importantly, these capacities appear in different combinations across occupations and can be developed through multiple curricular routes. A physics graduate may not be employed as a physicist, for example, but could nonetheless employ their quantitative and computational skills across a variety of jobs. Relatedly, a sociology graduate may not be trained for one specific occupation in the way that a nursing or engineering graduate is, but coursework in statistics, organizations, and social research may prepare that graduate for work in policy analysis, market research, or data-oriented organizational roles. Recent qualitative work similarly suggests that liberal arts graduates often apply specific degree-acquired skills, such as professional writing and research, to substantive work tasks — and sometimes to a greater extent than their peers in ‘practical’ majors (Moss-Pech, 2025). Taken together, these patterns demand a broader account of the relationship between what students learn in college and what they do at work.

A major limitation in prior studies is measurement. Existing approaches necessarily rely heavily on educational categories—most commonly field of study—as proxies for the skills workers possess (Gerber and Cheung, 2008; Altonji et al., 2012, 2014; Lovenheim and Smith, 2023). While educational categories are readily observed, they do not directly reveal the skill content of a graduate’s educational investments. Nor do they allow us to infer how closely a particular educational experience aligns with the task demands of the occupations graduates enter. A framework centered on transferable human capital therefore requires measuring educational content and work content in comparable terms. In the following section, I develop such a framework by representing graduates’ coursework and occupational tasks in a common task space and defining their alignment as *skill-task match*.

3 Skill-task match

I conceptualize college-acquired skills as portable investments that can be activated across a range of labor-market settings. College courses do not simply train students for specific occupations or discrete workplace tasks, but instead develop task-relevant capacities. These capacities are transferable across jobs that draw on similar underlying skills. For example, a physics course may prepare students for both a scientific role that requires detecting physical phenomena and reporting experimental results and a data science role that involves analyzing data and performing simulations. These jobs involve different substantive domains and different workplace tasks, but both draw on overlapping capacities in statistical reasoning, data analysis, and communicating quantitative evidence. In this sense, coursework in physics may provide preparation for both. Alignment between education and jobs can then be measured along a continuous spectrum: as the degree to which the capacities developed through a graduate’s coursework correspond to the capacities required by their job. I refer to this form of alignment as *skill-task match*.

3.1 Portable skill investments

For decades, sociological work has conceptualized occupations as groupings of similar tasks, skills, and responsibilities rather than as entirely discrete categories (Blau and Duncan, 1967; Hauser and Warren, 1997; Hout, 1984). Data sources, such as the Occupational Information Network (O*NET), verbatim text responses from respondents (Martín-Caughey, 2021), and, more recently, task taxonomies generated from job postings data (Han and Cheng, 2026) make it possible to characterize occupations in terms of their work content rather than only as nominal categories (Autor et al., 2008; Cheng et al., 2019; Deming, 2017; Goldin and Katz, 2008; Liu and Grusky, 2013). Characterizing jobs in terms of work content helps describe structures of inequality that broad occupational categories would obscure—for example, by explaining the growth in between-occupation wage inequality in recent decades (Liu and Grusky, 2013), and by characterizing intergenerational mobility more parsimoniously than broad labor-market categories (York et al., 2025).

Leveraging the work content of occupations enables researchers to take a *relational* perspective on the labor market. Occupations can be understood as more or less similar to one another depending on the tasks they require and the skills those tasks draw upon. Recent work formalizes this intuition by representing jobs in terms of co-occurring skill and task requirements, making it possible to measure similarity among labor-market positions rather than treating occupations as unrelated categories (Han and Cheng, 2026). This relational view also underlies theories of task-specific human

capital: workers can transfer skills to new occupations when the new occupation relies on similar tasks to the old one (Gathmann and Schönberg, 2010; Yamaguchi, 2012). Figure 1 shows this intuitively using a simplified two-dimensional representation of three target occupations and a physics graduate. In this representation, the angle between vectors captures similarity in task content. A data scientist, for example, may be expected to manipulate data with statistical software and identify business problems, while a physics-related scientific worker may be required to analyze data to detect physical phenomena and report experimental results.

Data-science and physics-related occupations therefore occupy nearby regions of the task space, whereas journalism sits farther away because it relies more heavily on interviewing, reporting, narrative writing, and public communication. Importantly, this proximity can arise even when the occupations do not share exact task requirements. Data-science and physics-related occupations may contain different concrete tasks in the O*NET data, but those tasks can nevertheless draw on related capacities: quantitative reasoning, statistical modeling, computational analysis, interpretation of empirical patterns, and communication of technical evidence. In this sense, the task space captures higher-order similarity in the underlying skill requirements of work, rather than only literal overlap in observed task labels.

If human capital is portable across occupations when occupations rely on similar tasks, then college-acquired skills should be portable across occupations that require similar skills. Yet in contrast to the detailed data infrastructures for measuring what workers do on the job, a comparable infrastructure for measuring what skills students acquire in college is far less developed. As a result, the large literature on the labor-market effects of higher education has often bypassed direct measures of skill acquisition, using field of study as the primary measure of educational content (Gerber and Cheung, 2008; Altonji et al., 2012, 2014; Lovenheim and Smith, 2023; Mountjoy, 2024). But field-of-study classifications do not, in themselves, reveal the skill content of education or its returns. Nor do they allow us to infer how closely a particular educational experience aligns with the task demands of the occupations graduates enter.

To address this gap, I apply the relational task perspective used to characterize occupations to the study of education. Rather than treating higher education primarily as a set of credentials or fields of study, I conceptualize educational investments in terms of the labor-market tasks for which they prepare students. This approach makes it possible to locate both educational experiences and occupations in a shared latent task space. In this way, I define *skill-task match* as the degree of proximity between a student’s course-derived skill portfolio and the task content of the occupation they enter. To construct this measure, I treat courses as the building blocks of students’ skill portfolios, inferring each student’s skills from their transcript trajectory, and comparing that portfolio to the task content of their job.

Figure 1 illustrates this logic. The black arrow represents a physics major. From a credential-to-occupation perspective, this graduate’s expected occupation is a Physicist; the graduate would be considered unmatched if they become a Data Scientist or Journalist. From a skill-task perspective, however, a physics major is located close not only to physicist jobs but also to data-science jobs. This is because the skillsets that a physics major imparts share task families with the data scientist: mathematical modeling, computational tools, statistical analysis, and quantitative reasoning.

This framework implies that common credential-to-occupation pathways and skill-task match are related but analytically distinct dimensions of education–work alignment. Table 1 shows a cross-tabulation of skill-task match with following a common credential-to-occupation pathway. Following a common credential-to-occupation pathway coincides with high skill-task alignment (top-left cell) when the expected occupational destination for a credential draws closely on the skills that credential imparts. This could occur, for example, in highly linked fields, where the strength of a field’s linkage could imply occupation-specific schooling that a graduate then leverages in their work (although high-linkage could also imply the presence of licensing regimes without a high degree of skill specificity) (DiPrete et al., 2017; Bol et al., 2019). Meanwhile, the lower-right cell represents the clearest case of mismatch or skill underutilization: graduates who enter less common destinations that do not draw strongly on their college-acquired skills.

The advantage of a skill-task match approach is that it separates the content of education–work alignment from pursuing a common credential-to-occupation pathway. The off-diagonal cells capture cases where skill-task alignment may differ from credential-to-occupation alignment. In the top right cell, individuals pursuing a common pathway may nevertheless fail to leverage their college skills at work (for example, in less linked fields of study). The lower left cell is the key case that a credential-to-occupation alignment approach obscures. A graduate may enter an occupation that is uncommon for their field while still using skills developed through coursework – precisely because those courses develop task-relevant capacities that transfer across occupations with related task demands. A relational task-space approach therefore expands the study of alignment beyond expected credential-to-occupation trajectories.

In addition to expecting skill-task match to vary substantially among individuals who both follow and do not follow a common credential-to-occupation pathway, I expect skill-task match to be associated with earnings, even net of educational inputs and net of a worker’s occupation. This core implication follows from task-specific human capital theory: workers should be more productive when the tasks required in their jobs draw on skills they have already developed (Gathmann and Schönberg, 2010; Yamaguchi, 2012; Poletaev and Robinson, 2008). Signaling models would also suggest an earnings payoff to skill-task match, though for different reasons: workers better matched to a job

Table 1: Conceptual relationship between common credential-to-occupation pathways and skill-task match.

	High skill-task match	Low skill-task match
Common credential match	Common pathway with content alignment: especially in fields with occupation-specific training.	Common pathway without strong content alignment: in fields with lower-linkage, or in high-linkage fields without occupation skill specificity.
Off-pathway	Transferable skill match: the worker departs from the common pathway but enters a job that uses transferable course-acquired skills.	Mismatch: neither the common pathway nor the measured task content aligns strongly with the worker’s education.

are able to communicate their fit for the job more clearly than workers who are not; this fit is rewarded by employers, even if match does not improve a worker’s productivity directly. College coursework may therefore generate returns outside expected pathways whenever it prepares students for task bundles that are valued elsewhere in the labor market.

At the same time, the returns to transferable skill alignment may be weaker or less immediate than the returns to common credential-to-occupation match. Occupation-specific pathways often make fit highly legible: credentials, professional categories, and common hiring pipelines tell employers what a worker has been trained to do (Spence, 1973; Arrow, 1973; Bills, 2003). They also may indicate the presence of licensing regimes that restrict supply for the occupation and bolster wages for those who enter the occupation. By contrast, transferable skills are less standardized and may be harder to verify before hire. A sociology major’s training in survey design, organizational analysis, or statistical reasoning may be useful in people analytics, but that fit is less categorically obvious than the fit between a nursing degree and nursing work. The absence of licensing and professional institutions may also lead to lower returns to alignment for those outside common credential-to-occupation pathways. As a result, broader skill-task match should predict earnings, but its effects may be smaller overall, and depend on whether workers and employers can identify, communicate, and validate the relevant skill overlap. This discussion leads to two hypotheses:

Hypothesis 1: Skill-task match will vary substantially within occupations,

majors, and expected match categories, especially among graduates outside common destinations.

Hypothesis 2: Graduates whose transcript-derived skill portfolios are more closely aligned with the task demands of their jobs will earn more than otherwise similar graduates, even after accounting for the overall economic return to their skill portfolios, credentials, and occupations.

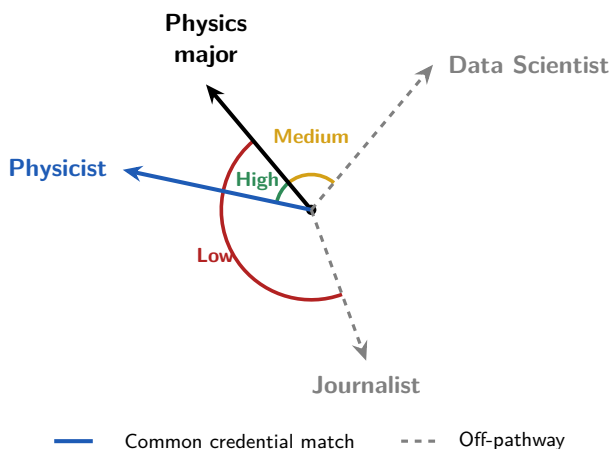


Figure 1: The “skill-task” space: Arrows depict the simulated “Physics Major” student (Black) and three target occupations as vectors emanating from a common origin. Line style encodes credential-to-occupation match (solid colored = common, dashed gray = off-pathway). Color of the arc and label at the origin encodes skill-task match (green = high, gold = mid, red = low).

3.2 Who gets to use their college skills? Disparities in matching and earnings inequalities

Sociologists have long treated occupational allocation as a central mechanism of labor-market inequality. It is well established that workers of color and those from lower socio-economic backgrounds remain disproportionately concentrated in lower-paying and lower-status occupations (Blau and Duncan, 1967; Breen and Jonsson, 2005; Mandel and Semyonov, 2016; Tomaskovic-Devey et al., 2006; Wingfield and Alston, 2013). Racial segregation across broad occupational groups, in particular, accounts for an important share of persistent Black–White earnings inequality in the United States. These disparities are partly rooted in unequal educational attainment, but they do not disappear among highly educated workers: racial and socio-economic differences in access

to lucrative occupations persist among college graduates, and contribute meaningfully to earnings gaps between racial groups and graduates from lower- and higher-income backgrounds (del Río and Alonso-Villar, 2015; Zhou and Pan, 2023; Witteveen and Attewell, 2017b; Laurison and Friedman, 2024).

But the jobs graduates obtain matter not only because occupations hold different positions in the hierarchy. They also differ in the extent to which they allow graduates to deploy the skills they acquired in college. Occupational allocation can therefore shape earnings not only by sorting workers into positions with different average levels of pay and status, but by determining whether their educational investments are either activated or underused. I hypothesize that skill transferability creates an additional margin of inequality in the school-to-work transition, beyond well-known racial and class sorting across the occupational hierarchy. Even graduates entering occupations of similar status may differ in whether they obtain work that aligns with their particular skill portfolios. While both workers and employers have incentives to form productive matches, job seekers conduct their searches under substantial informational and structural constraints (Parasurama et al., 2025) – constraints which limit the ability of minority job seekers to effectively match into skill-aligned work.

The first constraint is informational. It has long been noted that forming worker–job matches is difficult because both workers and employers possess incomplete information about one another (Akerlof, 1970; Spence, 1973; Granovetter, 1981). But access to the information needed to navigate this broader opportunity set is itself unequally distributed. Workers often have difficulty knowing in advance which available jobs best fit their abilities and preferences (Halaby, 1988), and this problem may be especially acute where there is no established credential-to-occupation pathway to guide search. Occupation-specific pathways provide relatively clear signals about plausible destinations, but outside these established pathways, the range of potentially suitable destinations is wider and the correspondence between schooling and work less obvious. Graduates must therefore discover which among many possible firms and occupations will actually make use of what they have learned. More advantaged graduates may be better positioned to overcome these informational frictions through social networks, institutional knowledge, and familiarity with employers and their selection processes. Privileged workers, for example, can draw on networks and strategic knowledge about elite firms to identify and obtain positions suited to their backgrounds and capabilities (Rivera, 2012). Transferability may therefore increase the value of precisely the kinds of labor-market information and social resources that are unequally distributed across graduates.

Informational frictions are compounded by discrimination. Hiring discrimination remains a persistent feature of the U.S. labor market, including for highly educated workers (Kline et al., 2022; Quillian et al., 2017). Employers may exclude minority

and lower-class applicants from jobs that would otherwise provide strong matches for their skills (Bertrand and Mullainathan, 2004; Pager et al., 2009; Rivera, 2012). In this sense, discrimination may not merely push disadvantaged graduates downward in the occupational hierarchy, but also prevent them from reaching particular firms and jobs in which their existing skills would be most productively deployed. Such demand-side exclusions can in turn reshape job-search behavior. Experiences or expectations of discrimination may alter which jobs workers consider worth pursuing in the first place (Barbulescu and Bidwell, 2013), especially when perceptions of discrimination are widespread (Goldsmith et al., 2004). In anticipation of discrimination by employers, minority graduates may therefore conduct broader and less targeted searches (Pager and Pedulla, 2015; Parasurama et al., 2025). Because applicants cannot know in advance which employers or occupations will be most discriminatory (cf. Heckman, 1998), casting a wider net increases the chance of receiving an offer and shortening the job search spell. But broader search may come at the cost of match quality: workers may accept jobs that are less closely aligned with their prior training. Recent work similarly suggests that repeated broad searches can accumulate over careers, producing less specialized employment histories and widening racial differences in labor-market trajectories (Parasurama et al., 2025).

Finally, unequal economic resources can constrain how long and how widely graduates are able to search for an appropriate match. Graduates from more affluent families may be better able to remain unemployed while searching, turn down poorly matched offers, relocate for desirable positions (Kogan et al., 2017), or absorb the costs of internships and other stepping-stones into particular careers (Laurison and Friedman, 2016; Friedman and Laurison, 2019). Less advantaged graduates may instead face pressure to accept the first viable offer available, even when it makes limited use of their skills. By restricting the set of employers and jobs that disadvantaged graduates can feasibly reach, these spatial and financial constraints may therefore undermine skill-task match even among workers with comparable educational preparation. This discussion therefore leads to the following hypothesis:

Hypothesis 3a: Conditional on pre-college preparation and college-acquired skill profiles, Black, Hispanic, and lower-income graduates will enter occupations with lower skill-task alignment than their more advantaged peers.

Because skill-task match itself carries an earnings premium, these disparities in access to skill-aligned work should also contribute to earnings inequality. Importantly, this contribution should be distinct from the earnings consequences of sorting across the occupational hierarchy:

Hypothesis 3b: Group disparities in skill-task alignment will contribute to group earnings disparities net of differences in where graduates are located in the occupational hierarchy.

4 Constructing skill-task match

The construction of skill-task match proceeds in three steps. First, I build an occupation-by-work-activity matrix from O*NET, representing each occupation as a bundle of task statements, and use a Singular Value Decomposition (SVD) to recover a lower-dimensional task space from this matrix. Second, I embed course descriptions and O*NET work-activity statements using a sentence embedding model, and use these embeddings to map each course onto the O*NET work activities to which it is semantically closest. Third, I aggregate course-level work-activity scores across each student’s transcript to construct a student-level skill profile, and project both student skill profiles and occupation task profiles into the same SVD-reduced task space. Skill-task match can then be measured as the cosine similarity between these two profiles.

I begin by constructing the occupational task space. The starting point is an occupation-skill matrix $\mathbf{M}$, where rows index occupations ($m = 873$ at the SOC-detailed level) and columns index O*NET work activities ($n = 2,087$). Work activities are standardized descriptors of job tasks compiled by the U.S. Department of Labor, providing information on the work activities required across occupations (for example, “analyze data to identify trends,” “develop statistical models,” or “coordinate project activities”). Unlike broader skill or occupation categories, work activities reflect the specific tasks workers perform on the job, and therefore provide a comparable representation of task content across occupations. Entry $M_{op} = 1$ if work activity p is associated with occupation o in the O*NET activity statements.

This raw matrix represents each occupation as a bundle of work activities and could, in principle, be used directly to calculate distances between occupations. However, because this representation is high-dimensional and sparse, occupations that require similar forms of work may not share many identical work activity indicators. For example, physicists and data scientists may both require quantitative reasoning, modeling, and analysis, but these similarities may be distributed across different activity statements. To capture this higher-order similarity between tasks and occupations, I apply a Singular Value Decomposition,

$$\mathbf{M} = \mathbf{U}\mathbf{\Sigma}\mathbf{V}^\top,$$

and retain the first $k = 24$ dimensions, selected using an elbow criterion on the singular-

value scree plot:

$$\mathbf{M}_k = \mathbf{U}_k \mathbf{\Sigma}_k \mathbf{V}_k^\top.$$

This decomposition embeds both occupations and work activities in a shared low-dimensional space. Occupations are represented by the rows of $\mathbf{U}_k \mathbf{\Sigma}_k$, while work activities are represented by the rows of $\mathbf{V}_k \mathbf{\Sigma}_k$. The resulting geometry captures higher-order task similarity: two occupations can be close not only because they share the same work activities, but because their work activities tend to co-occur with similar sets of other tasks across the occupational structure (see Han and Cheng, 2026). In this space, physicists and data scientists are more similar than physicists and journalists.

Table 2: Illustrative low-dimensional representation of the skill-task space. The first row gives a transcript-derived skill vector for a physics major, while subsequent rows give occupation task vectors based on O*NET work activities projected into a simplified three-dimensional SVD space. The final column reports the cosine similarity between the physics-major vector and each occupation vector. Entries are shown as binary 1/0 indicators for illustrative purposes only.

	Scientific / modeling domain	Applied data / business domain	Interviewing / writing domain	Skill-task match
Physics major	1	0	0	–
Physicist	1	0	0	1.00
Data Scientist	1	1	0	0.71
Journalist	0	0	1	0.00

The second step is to map each college course to its most semantically related work activities. To do this, I first embed each cleaned course paragraph (see next section for details) as well as the O*NET work activity statements into a vector space using a sentence embedding model. For each (course, work activity) pair (c, p) , I compute the cosine similarity between the course paragraph embedding $\mathbf{e}_c$ and the work activity descriptor embedding $\mathbf{s}_p$, resulting in a course–work activity similarity matrix $\mathbf{A}$ with approximately 290,000 rows (courses) and columns for each O*NET work activity. Each entry a_{cp} represents the semantic proximity between a course description and a

specific occupational activity, ranging from -1 (opposite meaning) to $+1$ (identical meaning). Because each course is plausibly unrelated to most work activities, for each course c , I retain only the top 5% of work activities by cosine similarity to the course description (the $q = .95$ threshold), setting all other similarities to 0. The results are not sensitive to the exact choice of threshold.

Finally, I compute a student’s skill profile via the simple sum of each work activity score across courses that the student takes. In alternative specifications, I weight each course by a student-specific term reflecting course intensity and performance, and apply an element-wise concave transformation which models diminishing marginal returns to repeated exposure to similar content. I also estimate a specification that controls for course dissimilarity, defined as the average pairwise dissimilarity across all courses a student takes, to test whether my measure of match may partly capture differences in curricular breadth. The results are not sensitive to these alternative modeling procedures (see Appendix D). I project each person’s skill profile as well as each task profile into the same 24-dimensional task space obtained via the Singular Value Decomposition. I use the resultant individual-level “Skill Profile” as a control in all the analyses that follow, in order to adjust for differences in detailed course investments even net of field of study.

Having constructed a course-derived skill vector for each student and an occupation-derived task vector for each job (in the same 24-dimensional task space), I measure skill-task match D_i as the cosine similarity between the two:

$$D_i := \text{sim}(Z_{1i}, Z_{2i}) = \cos \theta = \frac{Z_{1i} \cdot Z_{2i}}{\|Z_{1i}\| \|Z_{2i}\|}, \quad (1)$$

This score, which lies between -1 and 1 , captures the extent to which a graduate’s actual skill investments align with the tasks required in their job. A higher cosine similarity between these vectors indicates a stronger match between an individual’s educational background and their job’s tasks, while lower similarity reflects skill mismatch. This procedure is shown graphically in Figure 2, and illustrated for a physics major and three example occupations in Table 2 using a simplified three-dimensional version of the SVD-reduced task space.² Here, the physics profile overlaps most closely with the task vector for physicists, somewhat less with the task vector for data scientists, and least with the task vector for journalists. The resulting cosine similarities show how the same educational profile can be strongly aligned with a common occupational destination, partially aligned with an uncommon but task-adjacent occupation, and weakly aligned with an occupation whose work activities draw on a different set

²To aid intuition, Table 2 reports entries as binary 1/0 indicators; in the actual implementation, entries are continuous SVD coordinates (real-valued loadings on each latent task dimension), and the cosine similarity is computed from those continuous vectors.

of skills. I use the similarity scores estimated in Equation 1 as both an outcome and a treatment measure: as an outcome to examine variation in match as well as inequalities in match, and as a treatment to assess the economic effects of match.

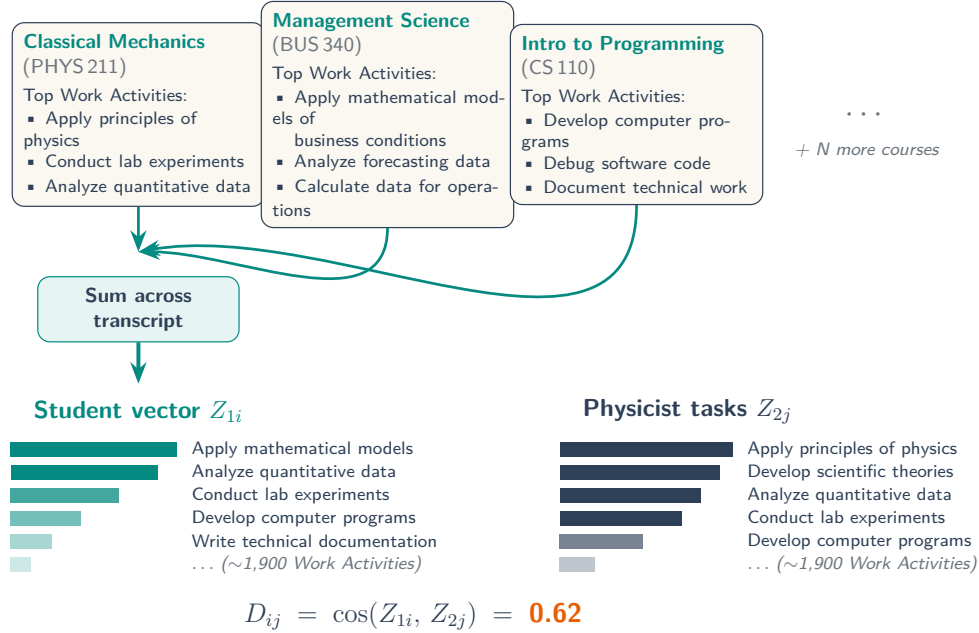


Figure 2: The “skill-task match” framework: each student’s transcript is transformed into a continuous, multidimensional skill profile by aggregating embedded course descriptions in the O*NET work activity task space; each post-college occupation is represented in the same task space using its O*NET work activities; and skill-task match is the cosine similarity between the two vectors. This approach captures forms of alignment between schooling and work that extend beyond binary credential-to-occupation linkages.

5 Data

I analyze skill-task match among BA graduates in the context of one U.S. state’s public 4-year college sector using a new anonymized dataset that links several sources of administrative data from that state.³

The first dataset is earnings data. These data come from the state’s Unemployment Insurance (UI) program, which requires nearly all employers to submit quarterly

³These datasets were linked to each other at the individual level by the state data agency, using social security number and identifying information.

reports on employee wages, hours, and occupations. These reports cover most private, local, and state public-sector workers. Notably, several groups—self-employed, federal employees, and active-duty military—are not covered. My main outcome is annualized earnings, measured in financial years 2023 and 2024, and adjusted to 2024 dollars using the CPI-U. Annual earnings are constructed by summing observed quarterly earnings, and workers with multiple jobs are assigned to the employer where they received the largest share of their compensation, provided they are observed at their primary employer for at least two quarters in any given financial year.⁴

Employers report occupation codes for each employee as part of their UI quarterly filings, using the standard six-digit Standard Occupational Classification (SOC). Coverage is partial: a non-trivial share of worker-quarter observations carry no SOC code, primarily because some employers either do not report occupation or report invalid codes. To assess whether this missingness is consequential, Appendix C compares the distribution of demographic covariates, prior academic preparation, and earnings between workers with and without SOC codes. Overall, the two groups are substantively similar on observables, though with some notable differences, particularly regarding employment: individuals with a reported SOC code work at larger employers that are more likely to be multi-establishment companies, and earn more. This pattern suggests that SOC-code availability reflects employer reporting capacity as well as worker characteristics: larger and more formalized firms may be better equipped to comply with mandatory occupation-reporting requirements.

Educational records come from the state’s postsecondary system as well as the high school system. Postsecondary files provide information on institution attended, GPA, credits, degree attainment, and postsecondary transcripts on courses taken in college. To ensure that education is complete before analyzing wages, I assign to each individual their highest credential attained in the *prior* financial year. If a student obtains multiple BA credentials in any given year, I attach to that student an in-state public degree if obtained, followed by a public or private 4-year out-of-state degree. If a student obtains multiple in-state BAs in a given year, I follow a rank ordering in which I assign students to the most selective degree obtained. Students who pursue graduate education are retained in the sample, but their degree attainment is coded to their highest BA institution. I also use enrollment records to identify a student’s credits, and cumulative GPA. High school transcript records provide GPA, school attended, and demographic characteristics (e.g., gender, race/ethnicity, disability status, and free and reduced-price lunch [FRPL] eligibility), as well as information on SAT and ACT scores provided by the College Board.

Finally, I develop skill representations of each course taken by students in the sam-

⁴I drop worker-firm observations with missing information on these attributes.

ple using publicly available course catalogs for each college-year in my sample. I obtain course catalogs between 2005 and 2022 for each of the five colleges via a combination of web-scraping (for colleges providing online catalogs) and Optical Character Recognition (OCR) procedures (for colleges providing PDF versions of their catalogs). Each course’s catalog text was pre-processed (using the procedure described in *Processing course catalog text* below) and then put into an embedding space yielding a numeric representation for each course. I then matched each course, as it appears in the administrative transcript data, to the numeric representation of the course as obtained from this embedding procedure. Because course names in the online catalogs differ slightly from course names in administrative transcript data, each catalog course was linked to the corresponding course in the administrative records via a combination of fuzzy string matching and hand-coding, restricting candidate matches to courses with the same CIP code and course number within the appropriate academic year. Figure A.1 in Appendix A shows that I successfully match 82% of all 354,806 unique courses taken by students over the period to a course catalog entry. I construct individual skill profiles by collapsing over the numeric representations of each course they took, as described in the previous section.

Processing course catalog text. I assemble course paragraph descriptions from these course catalogs using a multi-step procedure. Note that this procedure was performed on the public course catalog text; only the embedding representations of each cleaned catalog paragraph were matched to course names in the administrative records as described above. I first segment each course catalog description into sentences. Across the five institutions, 253,383 course descriptions with valid substantive text yield 741,961 sentences (that is, ≈ 2.9 sentences per description). A key implementation challenge was in accurately removing administrative and logistical content from course catalog descriptions. In preliminary tests, I found that boilerplate sentences (e.g., references to course prerequisites, scheduling, and assessment policies) artificially inflate descriptions’ similarity to education- and pedagogy-related work activities, biasing match measures upward for any course with a typical syllabus footer. I therefore filter sentences in two stages. First, a rule-based keyword filter discards sentences containing administrative or logistical language; this removes approximately 28% of sentences overall, and 6% of courses. Second, surviving sentences are submitted to an LLM for binary classification of skill relevance; only sentences classified as skill-relevant are retained (see Appendix A for details).⁵ Sentences passing both filters are concatenated in original order into a single paragraph per course; courses with fewer than five surviving words are dropped, and course-CIP descriptions are appended to provide additional content. The final output is one cleaned paragraph per course. Figure 3 illustrates

⁵Note that the LLM was run only on course catalog text; no individual-level data was passed through any LLM.

this cleaning process for a representative catalog entry. Overall, 82% of courses are successfully matched to the corresponding catalog entry; a further 11% of courses are lost because the keyword and LLM filters deem that the course catalog description contains no substantive (skill-related) content. Counts of courses remaining at each filter stage are reported in Appendix Figure A.1. On average, 77% of a student’s transcript courses are successfully matched and embedded after the keyword and LLM filters (median 81%; bottom decile below 57%). Although I successfully measure the majority of course investments for most students, it remains possible that the unmeasured courses reflect skill investments that are aligned or unaligned with their work tasks, which could bias my measure of match.⁶ To test for this possibility, I repeat the main analyses on the subsample of students for whom I am able to embed at least 80% of courses; the substantive results are unchanged (Figure D.1 in Appendix D).

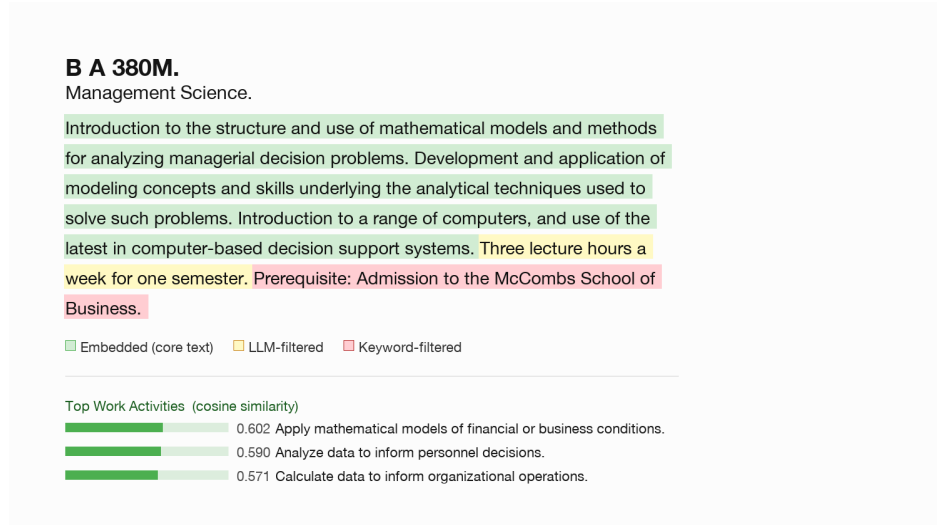


Figure 3: Illustration of the catalog-cleaning pipeline for a representative course. Raw catalog text is segmented into sentences; sentences containing administrative or logistical content are dropped by a rule-based keyword filter; remaining sentences are submitted to GPT-4o-mini for binary skill-relevance classification; surviving sentences are concatenated into a single cleaned paragraph used as input to the SBERT embedding step.

I also construct common credential-occupation match measures as a point of comparison with my text-based measure of match. To do this, I draw on the ACS 5-year sample 2019–2023. For each field-education-level cell, I identify matched occupations as follows. I first select the ten largest SOC 3-digit minor groups by weighted employment count among workers from that field and education level. I then compute an

⁶For example, I do not see the courses that students take in institutions outside the public state system.

actual-to-expected (A/E) ratio for each, defined as observed employment divided by the employment that would be expected if graduates from that field were distributed across occupations in the same proportions as all workers at that education level — that is, $A/E_{fgl} = n_{fgl} / (n_{gl} \cdot n_{fl} / n_l)$, following Bol et al. (2019). The two occupations with the highest A/E ratio among the top ten constitute the matched occupations for that cell. An individual is coded as matched ($= 1$) if they work in either of the two matched SOC minor groups for their field and education level, and unmatched ($= 0$) otherwise.

I focus my analyses on BA recipients who began their postsecondary careers at one of five (out of seven) public four-year colleges in the state, graduated with a BA between 2008 and 2022, and are observed with positive UI-covered earnings in the relevant period. This restriction is motivated by measurement feasibility. Constructing skill-task match requires a comprehensive transcript-derived skill profile, which in turn requires linking transcript courses to course-catalog text from the institutions students attended. Students who begin at two-year institutions and later transfer to four-year colleges typically attend a larger and more heterogeneous set of institutions, which would substantially expand the number of course catalogs that must be collected, processed, and linked. I therefore focus on four-year entrants in this analysis.⁷ Extending the analysis to two-year college entrants, associate-degree holders and certificate programs is an important avenue for future work.

Table 3 presents descriptive statistics for the core sample. Of the 74,811 workers in the overall sample of eligible BA graduates, 69% are observed with in-state earnings over the 2023-2024 period. Of these workers, 71% have employer-reported SOC codes. Workers with missing SOC codes for their primary employer-year observation are dropped from the matched-occupation analyses, since my measure of skill-task match requires both a course-derived skill profile and an occupation-derived task profile. Among workers with valid SOC codes, approximately 3,000 are excluded from the main analysis because their reported occupations do not have corresponding O*NET task ratings, which are required to construct occupation task profiles. This leaves me with a final analytic sample of $N = 32,694$.

5.1 Validating the measure

I validate the measure in two complementary ways. First, I examine whether skill-task match is higher along common credential-to-occupation pathways, and whether skill-task match varies in ways we would expect across different fields of study (see Appendix

⁷Students who transfer between four-year institutions are retained in the analytic sample. I do not explicitly drop BA graduates who also obtained associate degrees or other certificates, but in practice, students who begin their postsecondary studies at a 4-year college are unlikely to attain these credentials.

Table 3: Sample descriptives for BA completers.

Characteristic	Mean / proportion
Female	0.567
Black	0.029
Hispanic	0.070
White	0.667
Asian	0.194
FRPL (ever)	0.222
Final GPA	3.54
Has earnings	0.691
Has wage	0.690
Employer reported SOC code	0.491
Employer reported SOC code (cond. on earnings)	0.710
Log earnings	11.288
Log wage	3.898
N	74,811

B). Table B.2 shows that skill-task match is positively correlated with credential-based match ($r = .563$), indicating that graduates who enter conventional occupational pathways for their field also tend to have higher skill-task match scores. Consistent with this pattern, graduates classified as matched under the credential-based measure have substantially higher average skill-task match scores than those classified as unmatched (.609 versus .267). Figure B.1 further shows that average match scores vary in expected ways across fields of study. Fields conventionally characterized by weaker school-to-work linkages, such as the liberal arts, humanities, and social sciences (Bol et al., 2019), have the lowest average match scores, while fields with more structured pathways into work, such as health, engineering, and business, have the highest. These patterns provide an important validity check: although the skill-task measure is not defined in terms of common pathways, the measure recovers broad field-level differences that prior credential-based approaches would lead us to expect.

Second, I show that the measure responds correctly to the underlying course composition of a student’s profile. Specifically, I simulate physics-major students whose 30-course transcripts contain $N \in \{0, \dots, 30\}$ physics courses (randomly drawn from the corresponding portions of the course catalog) plus $30 - N$ non-STEM courses. For each N , I project the simulated transcript into the latent skill space and compute its cosine similarity to three target occupations: Physicists, Data Scientists, and Journalists. Figure 4 shows the result. As expected, similarity to Physicists rises sharply with N . Crucially, similarity to Data Scientists also rises with N —capturing the skill transferability the measure is designed to recover, since physics coursework develops

mathematical and computational capacities that data-scientist work requires—while similarity to Journalists declines.

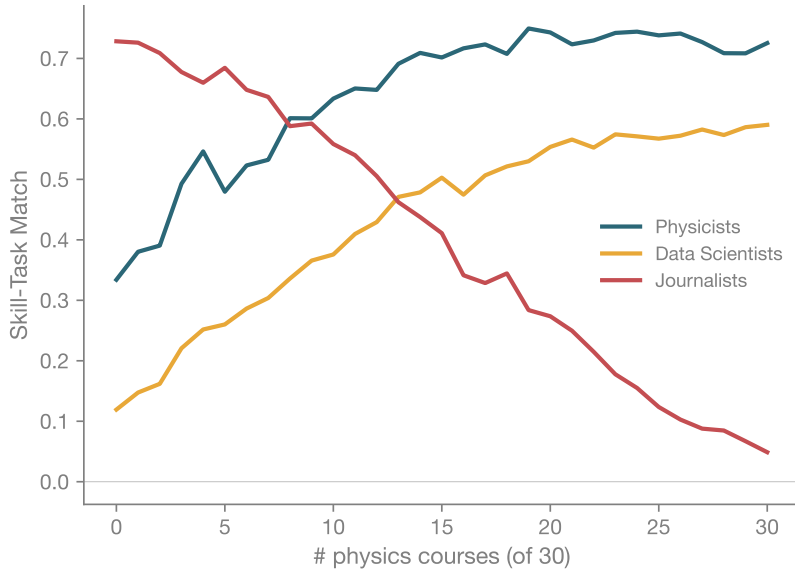


Figure 4: Skill-task match between simulated physics-major transcripts and three target occupations (Physicists, Data Scientists, Journalists), as a function of the number of physics courses (out of 30) on the transcript. Similarity is computed in the work-activity-based latent skill space.

6 Results

6.1 Skill-task match: variation and variance decomposition

I begin by characterizing the variation in skill-task match among individuals in my sample. Table B.1 shows that the average match score is .355; match scores have a standard deviation of .266, a 25th percentile of .125, and a 75th percentile of .585. We can gain an economic interpretation for these numbers in two ways. First, we can consider fixing a person’s occupation, and asking how much more coursework that person would need to take in a field associated with the occupation to obtain a one-standard-deviation increase in match. The validation simulation reported in Figure 4 shows that, for a fixed occupation, a one-standard-deviation increase in skill-task match corresponds to replacing approximately one third of an average student’s coursework with field-aligned courses, and moving a student from the 25th to the 75th percentile

of match corresponds to replacing approximately 80% of their coursework. Second, we can consider fixing a person’s coursework, and asking what sort of occupations would correspond to a change in match score. To give an illustration of this, for business majors who take more math-based courses, a one-standard-deviation increase in match corresponds to the difference between a math-oriented business major who becomes a data scientist, and a math-oriented business major who becomes a credit analyst.

What drives this variation? I examine the extent to which skill-task matching D_i varies within majors, occupations, schools, and a measure of making a common credential-to-occupation match. Specifically, I implement a variance decomposition of match as follows:

$$\text{Var}(D_i) = \underbrace{\text{Var}(\mathbb{E}[D_i|C_i])}_{\text{between-}C_i} + \underbrace{\mathbb{E}[\text{Var}(D_i|C_i)]}_{\text{within-}C_i},$$

where C_i denotes majors, occupations, or following a common match pathway. This decomposition allows me to quantify how much of the overall variation in skill-task congruence arises from systematic differences between categories (e.g., certain majors being better matched than others) versus within categories (e.g., students with the same major experiencing very different match outcomes).

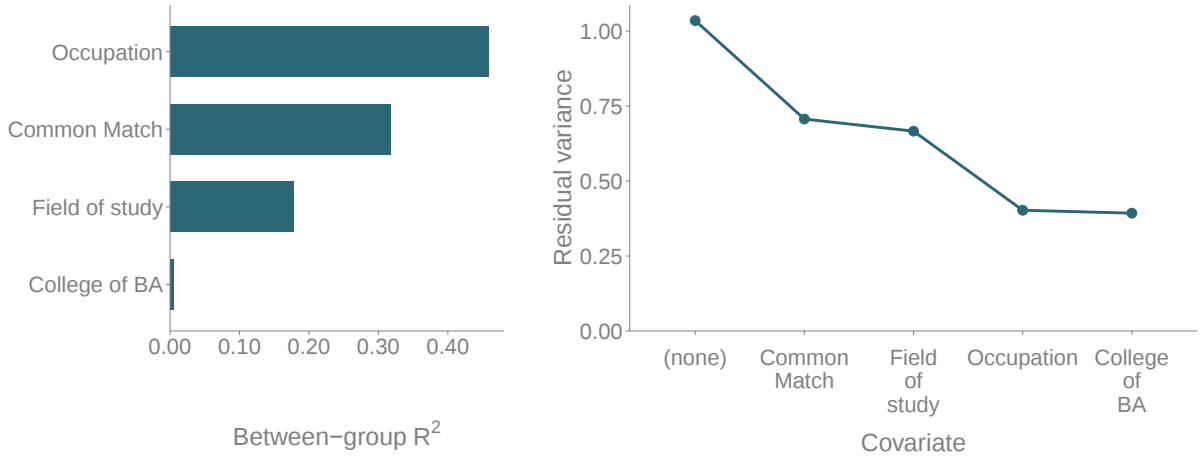


Figure 5: Left panel: variance decomposition of the skill-task match score. Bars show the share of total variance attributable to (i) field of study, (ii) occupation, (iii) BA college, and (iv) following a common credential-to-occupation pathway, each entered separately as a categorical predictor. Right panel: residual variance for skill-task match after the same predictors are added sequentially in a nested model.

The left panel of Figure 5 reports the results of this exercise. BA college explains

the lowest amount of variation: almost all of the variation in skill-task matching occurs among graduates of the same college, implying that specific colleges do not systematically sort their graduates into better-matched jobs than others. Field of study explains a modest portion of the total variation in skill-task matching, at 18%. Occupation explains the largest portion, at 46%, but a full half of the variation in match occurs within occupation groups: among individuals obtaining the same occupation, there are dramatic differences in the extent to which individuals are using their college-acquired skills. Importantly, common pathways between occupations and credentials contribute under one-third of the total variation in matching.

One limitation of the preceding variance decomposition is that it considers each factor independently. Yet, some of these factors are likely to be correlated: field of study, for example, partly shapes the occupations graduates enter, and one’s degree-granting institution shapes the fields of study available. To assess how much explanatory power each factor contributes net of the other factors, I estimate a series of nested regressions where I sequentially add common credential-occupation pathway, field of study, occupation, and BA college to a regression model of skill-task match. The results are shown in the right panel of Figure 5. The results are consistent with the variance decomposition exercise. Just under one third of the variation in skill-task match reflects whether an individual follows a common or uncommon pathway. Adding field of study to the regression model only decreases the residual variance by a small amount, whereas adding occupation decreases the variation considerably. In other words, some occupations tend to have systematically higher levels of alignment than others. Finally, adding BA college to the model does not explain additional variation in skill-task match. Collectively, field of study, occupation, BA college, and common match pathways account for 62% of the total variance in skill-task match. While the skill-task match measure therefore partly overlaps with these measures, skill-task match captures dimensions of the school-to-work transition that are independent of common credential-to-occupation linkages, providing support for Hypothesis 1.

6.2 Returns to skill-task match

Is skill-task match associated with earnings? To provide a descriptive test of this proposition, Figure 6 shows a binned scatterplot of log earnings against standardized skill-task match score, separately for graduates whose job is on a common credential-to-occupation pathway for their major and graduates whose job is off-pathway. The size of each bin is proportional to the number of graduates in each. I annotate several of the bins with the modal occupation in each.

The figure shows substantial variation in match, across both common and uncommon pathways. Examining just variation along the x-axis of each panel shows that,

among workers who do and do not make common matches, there is a wide range in who leverages their college-acquired skills on the job. This reflects the result in Figure 5, that most of the variation in skill-task match is within, rather than between, common-match categories. Further, this variation in match is associated with earnings. Individuals with higher scores of skill-task match generally have higher average earnings than those with lower scores.

Interestingly, the shape of the association between skill-task match and earnings is similar both for individuals who follow a common pathway and those who do not, steepest for those with the lowest match scores, and plateauing for those with the highest scores. The steeper gradient between matching and earnings among those with lower match scores may partly reflect occupation attainment: individuals with the lowest match scores are more likely to work in jobs involving routine sales, administrative, and clerical tasks – jobs which are well-known to often pay less. Occupational sorting also appears to explain some of the outlier points in the figure. Education majors working in the teaching profession are among the best matched workers following a common pathway, but are also relatively low paid. Business and social science majors working as financial analysts are poorly skill-task matched, but highly paid. At the same time, skill-task matching does not appear to be coterminous with occupation-based sorting. Individuals working in several generally high-paying occupations, but with lower skill-task match scores, earn less than those in the same occupations but with higher skill-task match scores. Nevertheless, to ensure I am not simply picking up an occupational sorting effect, in the earnings models presented below I include occupation fixed effects. This allows me to ask how skill-task alignment is associated with earnings among BA holders working in the same occupation, but who vary in their skill-job alignment.

To be sure, a large portion of this earnings premium may be driven by the characteristics of graduates who successfully make a match. Appendix Table B.3 shows that highly matched candidates are more likely to be White, have higher GPA and standardized test scores, and are less likely to have been on free and reduced-price lunches (FRPL) in high school. To isolate these characteristics from the role of match, I estimate the partial effect of match on earnings, conditional on skills and other covariates. Specifically, I fit the following model:

$$Y_i = \beta_1 D_i + \beta_2 X_i + \beta_3 Z_i + \alpha_i + \epsilon_i, \quad (2)$$

where Y_i is log annualized earnings, X_i denotes pre-college and college covariates (high-school GPA, SAT score, high school fixed effects, field of study, degree-granting institution, race, gender, free and reduced-price lunch status), and α_i are occupation fixed effects (6-digit SOC codes). I also control for Z_i , which denotes the student’s skill-embedding vector—a low-dimensional control for the student’s college skill investments.

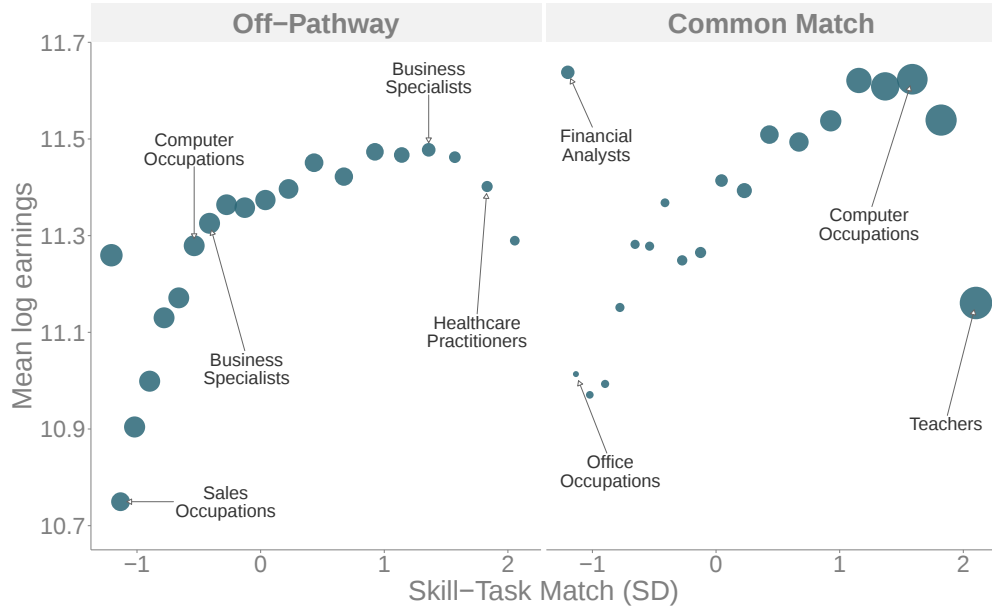


Figure 6: Binned scatterplot of log earnings on skill-task match score, separately for graduates on a common credential-to-occupation pathway and off the common pathway. Point size is proportional to the number of graduates in each bin.

I estimate parallel specifications that replace D_i with an indicator for making a common credential-to-occupation match, as well as specifications that include both skill-task match and common match jointly.

Figure 7 displays estimates of the earnings return to matching across each of these three specifications. The coefficients in blue show the estimated returns to common match (top row) and skill-task match (bottom row) separately. When estimated in separate models, both skill-task matching and matching along a common credential-to-occupation pathway yield a sizeable earnings return. In particular, a one-standard-deviation increase in skill-task match is associated with an earnings return of approximately 7%. Because of the rich set of controls, the skill-task match effect can be interpreted as the earnings difference between two BA holders with the same GPA, field of study, and from the same college who work in jobs with identical average earnings overall, but one of whom works in a job that is better aligned with their college skills.⁸ The earnings gain to skill-task match is also comparable in size to social skill complementarities reported by Deming (2017), who showed that complementarity be-

⁸Notably, my estimated payoff to matching along a credential-to-occupation pathway is smaller than the reported estimate in Bol et al. (2019). This attenuation is almost entirely driven by the inclusion of occupation fixed effects in my specification, which Bol et al. (2019) do not include in their models.

tween a worker’s social skills and social skill job requirements increases earnings by about 5%.

How distinct is the earnings return to skill-task match from the return to making a common pathway match? To address this, the coefficients in yellow present estimates of match earnings effects from a model which estimates returns to common match and skill-task match when both measures are placed in a model together. When both measures enter together, the estimated return to skill-task matching stays constant at 7%; by contrast, the estimated return to common matching attenuates to zero. That is to say, credential-to-occupation pathways do not carry an earnings premium over and above the actual alignment between graduates’ college-acquired skills and the tasks of their jobs. Instead, the premium associated with common pathway match appears to arise largely because these pathways are more likely to place graduates in work that uses the skills their education developed. Together, these results provide support for Hypothesis 2.

Table 4 reports the corresponding regression estimates across increasingly restrictive specifications. Without covariate adjustment, a one-standard-deviation increase in skill-task match is associated with a 0.139 point increase in log earnings. This association only declines slightly when pre-college controls are included, implying that pre-college achievement and resources are not strongly predictive of making a match. By contrast, including college controls attenuates the return to skill-task match by 26%, which primarily reflects the association between field of study and match as shown in Figure 5. Including occupation fixed effects attenuates the estimated match effect by another 30%, implying that part of the association between earnings and alignment comes from the fact that the most misaligned workers are found in jobs that are themselves low-paying (a point reflected in Figure 6). Finally, in Appendix E, I examine whether the earnings return to individual skill-task match reflects major-level alignment rather than student-specific course-taking, by comparing individual skill-task match to major-task match: the alignment between a graduate’s job tasks and the average skill profile implied by their field of study. The results show that the return to skill-task match is not reducible to the average skill profile associated with a graduate’s major, implying that students’ specific course trajectories capture worker–job fit beyond what can be inferred from field of study alone.

6.3 Skill utilization and earnings inequalities

Finally, I assess the role of skill-task match in driving demographic earnings gaps among college graduates. To do this, I first regress measures of match on three demographic characteristics: Hispanic (versus White), Black (versus White), and an indicator of class background (whether an individual ever received free and reduced-price lunches

Table 4: Earnings returns to Skill-Task Match across nested control specifications.

	M1	M2	M3	M4	M5	M6
Skill-Task Match	0.139 (0.004)	0.134 (0.004)	0.103 (0.004)	0.102 (0.004)	0.070 (0.006)	0.070 (0.007)
Common Match	—	—	—	—	—	0.001 (0.010)
Pre-college controls		✓	✓	✓	✓	✓
College / Field of Study controls			✓	✓	✓	✓
High-school fixed effects				✓	✓	✓
Occupation fixed effects					✓	✓
Common Match						✓

Notes: Standard errors in parentheses. **Pre-college controls** comprise student demographics (female, Black, Hispanic), GED status, high-school graduation year, indicators for ever experiencing homelessness, free/reduced-price lunch, special education, or English-language-learner status, high-school GPA, and the standardized SAT/ACT composite score. **College / Field of Study controls** comprise BA-college fixed effects, field of study, college GPA, and the student's course-profile embedding (capturing the average skill content of courses taken). **Occupation fixed effects** are 6-digit SOC codes. **Common Match** is an indicator equal to one if the student follows a common credential-to-occupation pathway and zero otherwise.

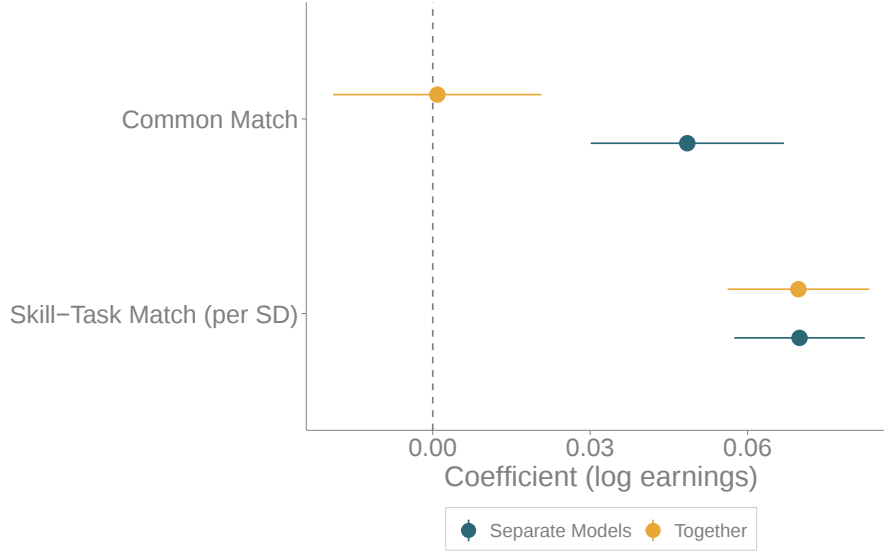


Figure 7: Earnings returns to skill-task match vs. common credential-to-occupation match across three model specifications. The key predictor is a one-standard-deviation increase in skill-task match (or a binary indicator of making a common match). Models control for pre-college characteristics (high-school GPA, SAT/ACT scores, race/ethnicity, gender, free and reduced-price lunch status), BA college fixed effects, field-of-study fixed effects, high-school fixed effects, occupation fixed effects (6-digit SOC codes), the low-dimensional transcript-derived skill profile, and the common credential-to-occupation match indicator. Point estimates with 95% confidence intervals.

in high school). In each specification, I include narrow controls for high school GPA, BA institution, field of study, as well as workers' skill embeddings. These gaps ask: among students with the same pre-college academic preparation, the same college and field of study, and the same latent skill profile, are minority or low-income students less likely to make a match? Similarly to the previous section, I conduct this exercise separately for skill-task matching as well as common pathway matching.

Figure 8 presents the results. The coefficient estimates in blue show estimated gaps in skill-task match, while the yellow point estimates show estimated gaps in making a common match, conditional on pre-college covariates, BA college and field of study, and individuals' skill profiles. Inequalities appear primarily off, rather than on, common match pathways. Demographic gaps in making a common match are small and (with the exception of the Black-White gap in match) statistically indistinguishable from zero. By contrast, under a skill-task match measure, the same comparisons recover large negative gaps for Black graduates, Hispanic graduates, and ever-FRPL graduates. The gap is particularly large between Black and White graduates, corresponding to

roughly a .19 SD gap in match. While these results alone do not provide evidence of mechanisms, they do provide some suggestive evidence that explicitly racialized mechanisms (such as differences in job search behaviors and hiring discrimination), rather than mechanisms due to financial resources alone (such as the importance of bridging funds), are most important in driving gaps in matching outcomes.^{9 10}

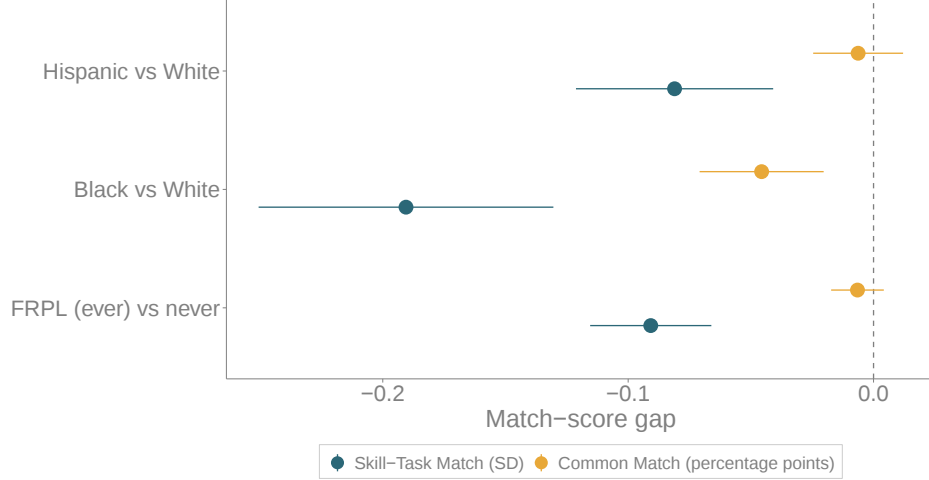


Figure 8: Demographic gaps in skill-task match versus common credential-to-occupation matching. Each panel reports the within-group difference (Black–White, Hispanic–White, FRPL-ever vs. never-FRPL) in mean skill-task match (blue) and in the probability of common match (yellow), conditional on pre-college covariates, college-level controls, field of study, and skill embeddings. 95% confidence intervals shown.

Figure 8 shows that minority and low-income students are less likely to obtain jobs that leverage their college skills, providing support for Hypothesis 3a. But this does not itself show how much these gaps matter for earnings gaps between these groups. To quantify the importance of these gaps for earnings disparities, I translate these match gaps into the earnings domain via a Kitagawa–Oaxaca–Blinder decomposition. Specifically, I decompose the total disparity in earnings between racial and family-income groups (among college graduates) into five components: pre-college characteristics, BA college, field of study and skill profiles, occupations and skill-task matching. Specifically, let X_i denote the column vector $X_i = [X_{i1}, X_{i2}, X_{i3}, X_{i4}, X_{i5}]'$, where X_{i1}

⁹The caveat here is that FRPL-status could act as too crude a proxy for family income. A continuous measure of family income could explain more of the Black-White gap in match, but such a measure was unavailable for my study.

¹⁰My finding of a small Black disadvantage in obtaining a common credential-to-occupation match aligns with that in Lu et al. (2025), who also find a small Black disadvantage in following a common occupational pathway using the Survey of Income and Program Participation.

denotes *pre-college characteristics* (high school GPA, English as a second language status, SAT/ACT score, and FRPL status for the racial earnings gaps), X_{i2} denotes *BA college*, X_{i3} denotes *field of study and skill embeddings*, X_{i4} denotes broad *occupational category*, and X_{i5} denotes *skill-task match*. Then, the total disparity between groups can be decomposed as (Blinder 1973; Oaxaca 1973)

$$\hat{\Delta} = \underbrace{\sum_{j=1}^5 \hat{\beta}_j^B (\bar{X}_{ij}^W - \bar{X}_{ij}^B)}_{\hat{\Delta}_{\text{Explained}}} + (\hat{\alpha}^W - \hat{\alpha}^B) + \underbrace{\sum_{j=1}^5 (\hat{\beta}_j^W - \hat{\beta}_j^B) \bar{X}_{ij}^W}_{\hat{\Delta}_{\text{Unexplained}}}. \quad (3)$$

where $\hat{\beta}_j^B$ and $\hat{\beta}_j^W$ denote the slope coefficients on X_{ij} for minority (e.g., Black) and majority (e.g., White) students, respectively, and $\bar{X}_{ij}^B$ and $\bar{X}_{ij}^W$ denote the corresponding group-mean values of X_{ij} . In this decomposition, $\hat{\Delta}_{\text{Explained}}$ consists of five components, each representing the contribution of each group of covariates described above. By contrast, $\hat{\Delta}_{\text{Unexplained}}$ captures group differences in the “returns” to each set of covariates (Jann, 2008).

Figure 9 shows the results of this exercise. For each demographic comparison, the point estimates decompose the raw log-earnings gap (red bars) into the portions explained by each component. Each component captures the direct contribution of the component, net of the subsequent components included in the decomposition. For example, the “Pre-College” component captures the direct contribution of pre-college resources, net of BA college, Field of Study / Skill Profiles, Occupation, and Skill-Task Match.

Pre-college characteristics explain a small (and in all cases, statistically insignificant) share of the total disparity between groups. This reflects the fact that, once BA college, field of study / skill profiles, occupation, and skill-task match are accounted for, race- and class-based differences in pre-college resources hold no additional explanatory power for the total earnings differences between these groups. For the FRPL-non-FRPL gap, BA college explains a slightly positive portion of the gap, whereas for the racial gaps in earnings, it explains a negative portion of the total gap. This results from the fact that, conditional on pre-college academic performance, Black and Hispanic students are in fact *more* likely to attend the Flagship school than their White peers—meaning that college-going actually compresses the gap between racial groups. Together, field of study and course profiles explain an important portion of the gaps (32% of the Black–White gap, 13% of the Hispanic–White gap, and 16% of the FRPL-non-FRPL gap). Occupational sorting also explains an important component of each gap (although its contribution is less precisely estimated for the Black–White gap) reflecting previous findings that racial minorities are disproportionately sorted into lower-paying occupations and are less likely to attain positions of authority and

managerial power (Tomaskovic-Devey et al., 2006; McTague et al., 2009; Kalev et al., 2006; Skaggs, 2009).

Skill-task match accounts for a non-trivial share of every gap. For the Black–White gap, its contribution is on par with that of field of study and skill investments: skill-task match accounts for 25% of the Black–White gap in earnings among college graduates. The portion explained by match is smaller, but still non-trivial, for the Hispanic–White and FRPL–non-FRPL earnings gaps, capturing 5% of the Hispanic–White gap and 8% of the FRPL–non-FRPL gap. Importantly, the portion explained by skill-task alignment is *net of* broad occupational category, which implies that race- and class-based sorting into skill-aligned work is an inequality-generating mechanism distinct from simply sorting into higher-paying work (Hypothesis 3b).

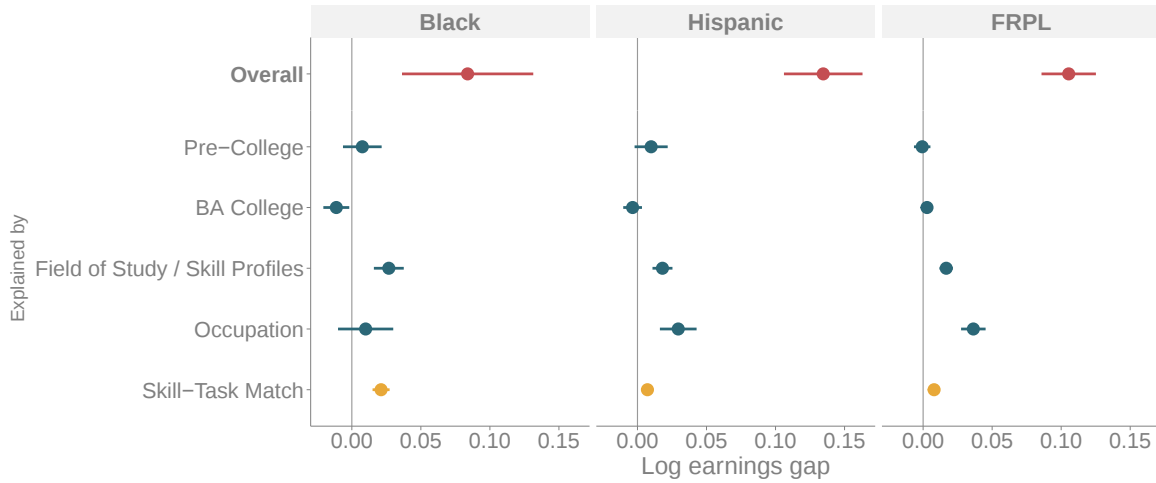


Figure 9: Kitagawa–Oaxaca–Blinder decomposition of the log-earnings gap, by group (Black–White, Hispanic–White, FRPL vs. never-FRPL). Within each panel, the top point (red) reports the unconditional log-earnings gap; subsequent points (blue) report the share of the gap attributable to covariate blocks (pre-college characteristics, BA college, field of study and course profiles, occupation); the bottom bar (gold) reports the share attributable to skill-task match. Each block’s contribution is incremental over the preceding blocks. Skill profiles are the low-dimensional work-activity representation of students’ transcripts. 95% confidence intervals from $B = 200$ bootstrap replicates.

7 Discussion

The transition from college to work is often understood in terms of formal categories: credentials, occupations, and the pathways that link them. This paper shows that

these categories miss an important dimension of labor-market alignment: the transfer of college-acquired skills across occupational boundaries. I measure this alignment directly by developing a novel approach to characterizing the skill content of education and applying it to a new linked dataset of four-year college graduates in a U.S. state.

The findings show that this form of alignment varies substantially and is consequential. Graduates vary greatly in whether they enter jobs that draw on what they learned in college; that variation persists within majors, occupations, and common credential-to-occupation pathways, so broad educational and occupational categories obscure meaningful heterogeneity in worker–job fit. The earnings return to skill utilization also extends beyond occupation-specific training: college-acquired skills are portable and rewarded across multiple labor-market destinations, and play an important role in explaining disparities among college graduates – especially between Black and White BA holders. The size of the earnings premium to skill-task match is notable precisely because it is not backed by the same institutions that organize conventional school-to-work linkages. Unlike tightly linked pathways, which may be reinforced by licensing regimes, professional categories, and recognizable hiring pipelines, transferable skill alignment is less credentialized and less immediately legible to employers. Yet the results show that this less formalized form of alignment nevertheless carries substantial economic value. The findings also broaden prevailing accounts of occupational allocation. Existing research emphasizes the sorting of workers into positions carrying different levels of pay and status. The results show that allocation also has a relational dimension. Because access to well-aligned pairings is stratified by race and class, unequal occupational allocation operates through skill realization as well as sorting across the occupational hierarchy.

There are several important limitations to these analyses. First, the data provide rare granularity into the relationship between schooling and employment, but are limited to one state’s public college system and labor market. Patterns of match may differ in private colleges, in colleges with stronger liberal-arts orientations, and in the two-year sector, where vocational programs may map more directly onto common occupational pathways. Second, the estimated earnings returns exclude graduates whose employment is not observed in the linked data. In particular, workers who obtain out-of-state employment are not covered by in-state UI earnings records. To the extent that individuals migrate out of state to pursue better-matched employment, this could attenuate my estimated earnings returns to match, but it is difficult to estimate the extent of the bias. Relatedly, I am unable to examine match for employers that do not report occupations for their workers. These employers are disproportionately smaller firms (Appendix C), and common credential-to-occupation pathways may be especially limited in precisely these settings.

Third, I use occupations rather than jobs to infer work tasks and to determine skill-

task match. Yet, such an approach could miscategorize individuals due to heterogeneity in task requirements across jobs within occupational categories (Martín-Caughey, 2021; Han and Cheng, 2026). A worker who appears mismatched to the overall task profile of an occupation may be well matched to a specialized position within that occupation, and vice versa. For example, a statistics major obtaining work as a biology researcher that requires expertise in biostatistics is highly skill-task aligned, even though my measure of skill-task match — at the level of the occupation — would mask this alignment. Future work linking transcripts to job postings, resumes, or employer-side task descriptions could distinguish within-job skill diversity from between-job heterogeneity inside occupations. Finally, my framework does not provide causal identification. While I control for a large set of college and pre-college covariates in my models, unobserved motivation, networks and family resources could explain the match premium that I document. Notwithstanding these limitations, this study provides the first granular view of the link between schooling and employment and of the transferability of higher education skills, and offers a first step toward understanding the earnings consequences of these dynamics.

8 Conclusion

Recent debates over higher education accountability increasingly ask whether programs deliver labor-market value. Accreditation reform has become a key site of this debate, with federal policy discussions placing renewed emphasis on whether institutions provide high-quality programs. These debates often presume that the labor-market value of higher education can be judged by whether programs prepare students for specific jobs. This paper suggests a more complicated picture. A large part of higher education’s value may lie not in narrow occupational preparation, but in providing students with transferable skills that can be mobilized and rewarded across multiple job domains. Skill specificity is one component of the return to higher education; judging programs only by single-occupation placement risks undervaluing broad forms of undergraduate training.

At the same time, showing that skills are highly transferable across economic domains also reveals an important dimension of inequality. Graduates do not benefit equally from the skills they acquire in college. One advantage of tightly defined credential-to-occupation linkages is that large numbers of graduates with these credentials are able to match into skill-utilizing jobs. Even if broader skill alignment is an important component of earnings, broader skill utilization and transferability necessarily come with higher labor market frictions. Many graduates in lower-linked fields may simply not know what labor market destinations would constitute a good match.

These frictions are compounded by resource disadvantages of lower-income groups, group differences in job search behavior, and hiring discrimination. Efforts to reduce postsecondary earnings gaps are therefore likely to be incomplete if they focus only on access to college, choice of major, or course-taking, without attention to the process through which graduates find skill-aligned work. Colleges, in particular, may have a role to play beyond the classroom: strengthening career services, alumni networks, and employer partnerships can help ensure that skills students acquire translate into better labor-market outcomes. At the same time, structural constraints—including family resources and bridging funds—limit what colleges can achieve without broader reform.

The framework developed here also opens broader questions about how skill investments matter over the life course. Over time, skill alignment is likely to be dynamic. Workers accumulate skills through jobs as well as through schooling (Gibbons and Waldman, 2004; Kalleberg and Mouw, 2018), and the relevance of those skills changes as technology, firms, and occupations evolve. Future research could therefore conceptualize match not only as the relationship between a graduate’s transcript and job, but as the relationship between a worker’s accumulated skill history and the changing task structure of available work.

Such a dynamic perspective could also recast skill-task match as a way to understand labor-market resilience. Job displacement may be less costly for workers whose skills align with multiple possible labor-market destinations, and more costly for workers whose skills are tied to a narrow set of tasks, firms, or occupational pathways. A task-based measure of match could therefore help identify not only who is well matched in their current job, but also who has viable alternative matches elsewhere in the labor market. In this sense, skill-task match provides a framework for understanding the capacity of workers to adapt as jobs, technologies, and task demands change.

Appendices

A Course-to-work-activity pipeline funnel

Here, I describe in more detail certain parts of the catalog processing that produces each student’s skill profile. The processing proceeds in four stages: (1) catalog ingestion (raw course descriptions scraped from each college’s online catalog); (2) administrative-record matching (catalog course $\rightarrow$ transcript course via exact code match, then fuzzy title match, then AI-assisted verification); (3) keyword filter (rule-based removal of administrative/logistical sentences); (4) LLM classification for skill-relevant sentences. Note that only course catalog text was passed through an LLM to tag for skill-relevant sentences in the course description; no student-level data was ever submitted to an LLM. I describe stages 3 and 4 in more detail below.

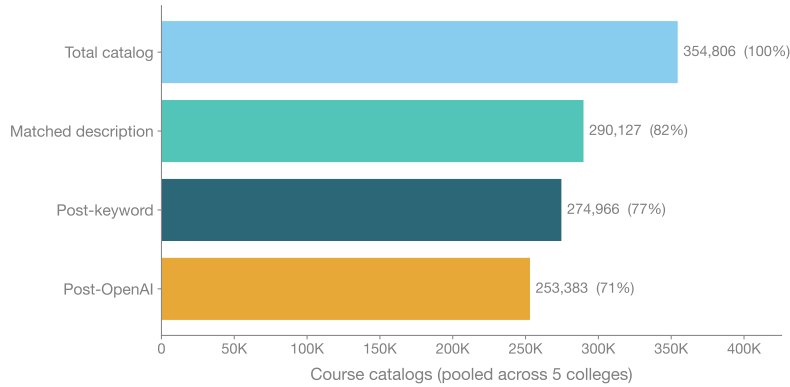


Figure A.1: Course-to-work-activity pipeline funnel: share of courses surviving each filter stage, pooled across all five colleges. Bars report cumulative retention from catalog ingestion through final work activity scoring.

A.1 Stage 3: keyword filter

The keyword filter drops any sentence containing one of the following terms (case-insensitive regex match, pipe-separated):

```
prerequisite|offered:|course details|myplan|instructors|offered|in  
a sequence|open only to|no initial knowledge|credit|concurrent  
enrollment|no auditors|required|meets professional  
standards|mandatory|restricted|sequence|continuation  
of|majoring|study abroad|for students|high school
```

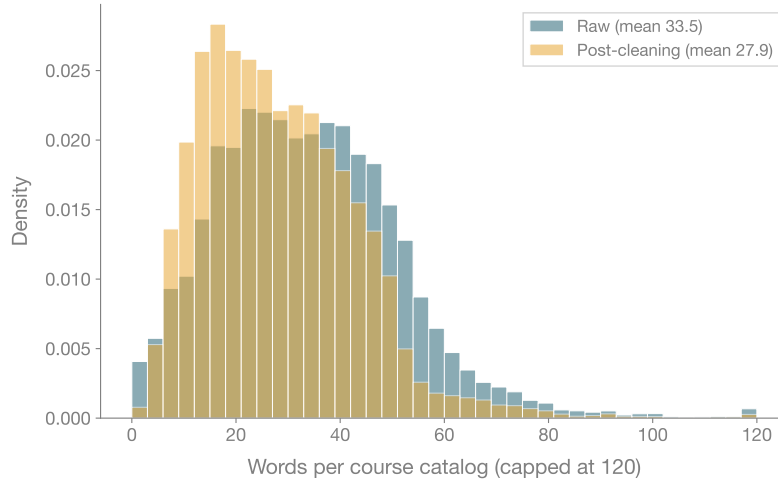


Figure A.2: Distribution of cleaned-paragraph word counts per course, pooled across all five colleges. The vertical line marks the 5-word minimum-length cutoff used to drop courses with insufficient surviving content.

The list was developed iteratively on a calibration sample to identify boilerplate sentences (prerequisites, scheduling, enrollment policies) that systematically inflate similarity to education- and pedagogy-related work activities.

A.2 Stage 4: LLM skill-relevance classifier

Surviving sentences are submitted to OpenAI’s GPT-4o-mini for a binary skill-relevance classification. The system prompt is:

You are an internal tool that classifies individual sentences extracted from university course descriptions (syllabi / catalog entries).

For each input sentence, determine **skill_relevance**: does it describe skill-relevant content (course topics, methods, concepts, competencies) or logistics/prerequisites (enrollment info, credits, scheduling, prerequisites)?

- 1 = skill-relevant course content
- 0 = logistics, prerequisites, or administrative information

Rules: Return EXACTLY one digit (**skill_relevance**); if ambiguous, lean toward 1 (keep it).

Example sentences classified as not skill-relevant. The following ten illustrative sentences received `skill_relevance` = 0 and were dropped before embedding:

1. “Prerequisite: MATH 124 or equivalent placement.”
2. “This course is offered every fall and spring quarter.”
3. “Students must obtain instructor permission before registering.”
4. “Final grade is based on three midterm exams (60%) and a final paper (40%).”
5. “Counts as a writing-intensive (W) credit.”
6. “Class meets Tuesdays and Thursdays from 1:30 to 3:20 PM.”
7. “Offered jointly with the Department of Economics.”
8. “This course satisfies the quantitative reasoning (QSR) general education requirement.”
9. “Restricted to students majoring in Computer Science with junior or senior standing.”
10. “Course materials provided through Canvas; no required textbook.”

B Additional sample descriptives

Table B.1: Distribution of the raw skill–task match score for BA completers. Columns report the sample size, mean, standard deviation, and key percentiles. The raw match score is the cosine similarity between a student’s transcript-derived work-activity vector and their occupation’s work-activity vector, prior to within-occupation standardization used in the regression analyses.

Sample	N	Mean	SD	p10	p25	p50	p75	p90
BA completers	32,694	0.355	0.266	0.031	0.125	0.300	0.585	0.757

Table B.2: Skill-Task Match by Common Match indicator.

	N	Mean	SD	p25	p50	p75
Common Match	8,438	0.609	0.228	0.485	0.671	0.779
Uncommon Match	24,256	0.267	0.216	0.092	0.218	0.400
Correlation (Common Match, Skill-Task Match): 0.563						

Table B.3: Sample descriptives by quartile of skill–task match. Columns compare students in the bottom and top quartiles of the (within-sample) Skill-Task Match distribution. Rows report demographic composition and pre-college academic preparation. Cell entries are percentages for categorical variables and means for continuous variables.

Characteristic	Quartile of Skill-Task Match	
	Bottom quartile	Top quartile
Female	56.5%	53.8%
Black	3.3%	1.4%
Hispanic	6.4%	5.1%
White	68.4%	70.7%
Asian	17.8%	18.2%
FRPL (ever)	20.6%	16.0%
Final GPA	3.50	3.60
Standardized test score	0.12	0.33

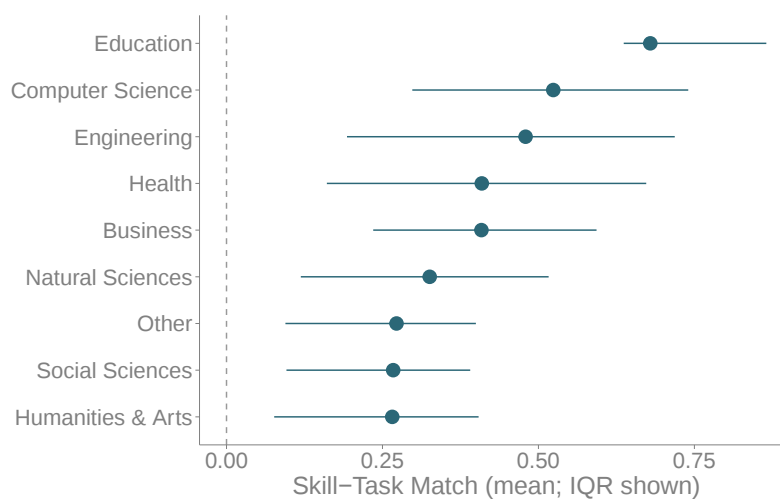


Figure B.1: Mean skill-task match score, by field of study (BA completers). Fields with weaker school-to-work linkages (liberal arts, humanities, social sciences) have the lowest mean match scores; fields with more structured pathways (health, engineering, business) have the highest. “Other” bundles low-headcount fields: architecture, protective services, family and consumer sciences, legal studies, other technical, communications and journalism, agriculture, public administration, and professional/culinary services.

C Sample balance

Table C.1 compares BA completers with positive UI-covered earnings to those with no observed earnings. The two columns are substantively similar on demographic and pre-college covariates; the No-earnings group has zero employer-reported SOC coverage by construction. Table C.2 similarly compares workers with and without an employer-reported SOC code among those with positive earnings.

Table C.1: Covariate balance for core sample with versus without UI earnings.

	Has earnings	No earnings
Female	56.4%	57.2%
Black	2.9%	3.0%
Hispanic	7.7%	5.4%
White	65.9%	68.5%
FRPL (ever)	23.2%	20.1%
Final HS GPA (mean, SD)	3.53 (0.43)	3.56 (0.44)
Standardized test score	0.15	0.32
Employer reported SOC code	71.0%	0.0%

Table C.2: Covariate balance for core sample with observed earnings, with and without SOC code.

	Has employer-reported SOC	No employer-reported SOC
<i>Student characteristics</i>		
Female	55.0%	59.9%
Black	2.5%	3.8%
Hispanic	6.1%	11.5%
White	68.6%	59.3%
FRPL (ever)	18.4%	34.8%
Final HS GPA (mean, SD)	3.54 (0.43)	3.52 (0.43)
Standardized test score	0.20	0.04
<i>Employer characteristics</i>		
Mean log earnings	11.31	11.23
Employer 10,000+ employees	26.7%	20.3%

D Robustness analyses

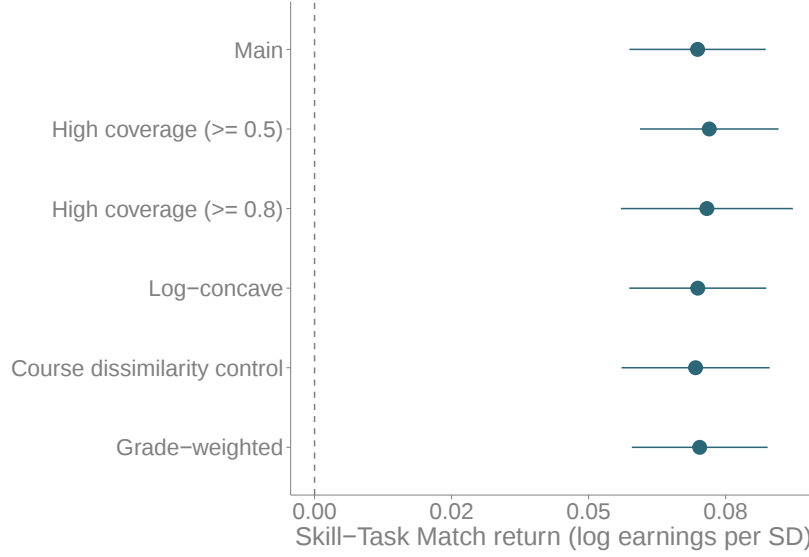


Figure D.1: Robustness of the main specification to alternative modeling choices. Restricting the sample to high-coverage students (those for whom at least 80% of transcript courses survive the keyword and LLM filters and are successfully embedded against the work-activity space) addresses the concern that systematic missingness in course-text content could be biasing the main results. Replacing the linear sum aggregation of work-activity scores with a concave $\log(1 + x)$ transformation models diminishing marginal returns to repeated exposure to similar course content—e.g., the marginal informational value of a fifth quantitative course is smaller than that of a second. In both robustness checks, the point estimates and 95% confidence intervals show that the earnings return to skill-task match is essentially unchanged.

E A major-level benchmark

The preceding argument treats transcript-derived skill-task match as a measure of the alignment between what graduates learned in college and the tasks they perform at work. Yet graduates enter the labor market not only with skills, but also with credentials. Declared majors organize course-taking and operate as socially recognized categories that help employers infer capacities, expertise, and appropriate job matches (Spence, 1973; Arrow, 1973; Bills, 2003; Rivera, 2011). As a result, graduates with high transcript-derived skill-task match may also hold jobs that align with the average skill profile associated with their major. If earnings rewards operate mainly through this broader category-level alignment, then the apparent earnings association with individual skill-task match may reflect major-implied alignment rather than the specific content of students' transcript trajectories.

I assess this possibility by comparing individual skill-task match to a parallel measure of major-task match: the alignment between the task content of the graduate's job and the average transcript-derived skill profile associated with the graduate's major. This benchmark is theoretically important because majors make some forms of worker-job fit more legible than others. Categories help audiences interpret actors and determine which criteria should be used to evaluate them (Zuckerman, 1999; Hannan et al., 2007; Leung, 2014). In labor markets, a job-seeker's major can perform this classificatory work by making some claims of fit more credible than others (Freidson, 1994; Weeden, 2002). From this perspective, the earnings association between individual skill-task match and earnings should be substantially reduced once major-task match is taken into account.¹¹

At the same time, majors are useful but incomplete summaries of skill acquisition. The U.S. credential landscape is fragmented, and school-to-work linkages are comparatively weak relative to coordinated market economies (Hall and Soskice, 2001). Students within the same major often complete different course sequences, while students in different majors may acquire overlapping skills through electives, general-education requirements, minors, projects, and cross-field coursework (Han et al., 2023; Kizilcec et al., 2023). Course-level investments may also become visible beyond the formal credential title: graduates can highlight specific courses, tools, or portfolios in resumes and interviews (Kreisman et al., 2021), and employers may use skill assessments or early

¹¹Empirically, Manuel (2024) measures the similarity between the skill requirements of an individual's occupation and the knowledge requirements of occupations closely associated with that individual's field of study. That paper finds an important payoff to major-occupation knowledge alignment, but cannot determine whether this alignment operates net of specific educational investments. The present analysis builds on this logic by asking whether the association between individual skill-task match and earnings is attenuated after accounting for major-task match, measured as the alignment between a graduate's job and the average transcript-derived skill profile associated with the graduate's major.

job performance to observe competencies that degrees and majors do not fully capture (Piopiunik et al., 2020; Sigelman et al., 2024). From both signaling and human-capital perspectives, then, transcript-derived skills may provide information about worker–job fit that broad credentials do not fully capture (Spence, 1973; Arrow, 1973; Stiglitz, 1975; Bills, 2003; Becker, 1962).

I therefore treat major-task match as a category-level benchmark for evaluating the interpretation of individual skill-task match. If the earnings association between individual skill-task match and earnings is substantially attenuated after accounting for major-task match, this would suggest that the observed association primarily reflects major-implied alignment. If the association remains substantial, this suggests that transcript-derived skill portfolios contain labor-market information beyond the implied skill profile of field of study.

To assess this possibility, I compare individual skill-task match to major-task match: the alignment between a graduate’s job tasks and the average skill profile implied by their field of study. I measure major-task match using national course-skill profiles from Sabet et al. (2024), who infer major-level skill profiles from more than three million course syllabi at nearly three thousand U.S. higher-education institutions. For each graduate, I assign the national skill profile corresponding to their field of study and compute its cosine similarity to the task profile of their occupation. This score captures how closely the graduate’s job aligns with the broader skill profile implied by their major, rather than with their own transcript trajectory. I then estimate earnings models that include individual skill-task match and major-task match separately and jointly.

Table E.1: Earnings returns to skill–task match versus a major-level skill profile constructed from national course-syllabus data (Sabet et al., 2024). Model 1 includes Average Major Match only, Model 2 includes Skill-Task Match only, and Model 3 includes both jointly. Standard errors are shown in parentheses. All specifications include pre-college controls, college-level controls, high-school fixed effects, and occupation fixed effects.

	Model 1	Model 2	Model 3
Skill-Task Match	–	0.070 (0.006)	0.071 (0.007)
Average Major Match	0.022 (0.004)	–	-0.002 (0.005)

Table E.1 reports the results. Major-task match is positively associated with earnings when entered in a model without skill-task match (conferring a .022 log increase in earnings per SD increase), suggesting that alignment between job tasks and the

average skill profile associated with a graduate’s field of study is rewarded in the labor market. However, once major-task match is entered in a model alongside skill-task alignment, the coefficient on individual skill-task match remains substantial, while the coefficient on major-task match attenuates to zero. This pattern indicates that the earnings return to skill-task alignment is not reducible to the skill profile implied by field of study. Instead, transcript-derived skill portfolios contain labor-market-relevant information beyond what can be inferred from a graduate’s major.

I also construct a major-level measure of match internally from the course catalogs in my data: for each major, I average the work-activity scores across all courses associated with that major, then compute the cosine similarity between this major-level skill profile and the task profile of the graduate’s occupation. These results, presented in Table E.2, are consistent with those in Table E.1.

Table E.2: Earnings returns to skill–task match versus a major-level skill profile constructed by averaging work-activity scores across all courses associated with each major in the catalog data. Model 1 includes Average Major Match only, Model 2 includes Skill-Task Match only, and Model 3 includes both jointly. Standard errors are shown in parentheses. All specifications include pre-college controls, college-level controls, high-school fixed effects, and occupation fixed effects.

	Model 1	Model 2	Model 3
Skill-Task Match	–	0.069 (0.006)	0.074 (0.008)
Average Major Match	0.036 (0.006)	–	-0.007 (0.007)

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